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### **ORIGINAL ARTICLE /** ARTIGO ORIGINAL

# Vigitel – Aracaju, Sergipe, 2008: the effects of post-stratification adjustments in correcting biases due to the small amount of households with a landline telephone

Vigitel – Aracaju, Sergipe, 2008: efeitos da pós-estratificação na correção de vícios decorrentes da baixa cobertura de domicílios com telefone fixo

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**ABSTRACT:** *Introduction:* The Telephone Survey has several attractions, such as low operational costs and speed the process, when compared to the Household Survey. However, the exclusion of house holds without a landline phone canpose a serious question of the valid it yof the estimates. *Objective:* Evaluating the use of post stratification adjustments to correct the potential biasdue to low coverage of landline phonein the results published in the Vigitel system. *Methods:* We compared the prevalence obtained by the Household Survey and the Vigitel, in Aracaju, Sergipe, where 49% of households had a telephone, calculating the differences measured between the surveys thatused the square root mean square error as a measure of accuracy of the estimate. *Results:* The Household Survey showed ten variables associated with landline phone. From this set of variables, post stratification weights corrected the potential bias of the consumption of beans in five or more days per week, vegetables intake regularly and recommended self-assessment of health as bad and morbidity of diabetes, while the weights partially eliminated the bias of the prevalence of asthma, possession of health insurance coverage and prevention of breast cancer in women aged 50 to 59 years. *Conclusions:* In order to reduce the potential bias in the results published by Vigitel system, in areas with low telephone coverage, it becomes necessary to use alternative weighting procedures and selection strategy of external variables for construction of post-stratification weights.

Keywords: Data collection. Interview. Population surveys. Selection Bias.

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**RESUMO:** Introdução: O Inquérito por Telefone, quando comparado ao Inquérito Domiciliar, apresenta vários atrativos, em especial baixo custo operacional e rapidez do processo de divulgação de resultados. No entanto, a exclusão de domicílios sem telefone fixo pode apresentar problemas na validade externa das estimativas. Objetivo: Avaliar os efeitos do uso de pós-estratificação para corrigir vícios decorrentes da baixa cobertura de domicílios com telefone nos resultados divulgados no sistema Vigitel. Métodos: Comparar resultados levantados no Inquérito Domiciliar com Vigitel no município de Aracaju, Sergipe, cuja cobertura era de 49% de telefonia fixa. O vício do Vigitel foi expresso pela diferença entre as prevalências do Vigitel e do Inquérito Domiciliar, sendo calculada a raiz quadrada do erro quadrático médio como medida de acurácia da estimativa. Resultados: O Inquérito Domiciliar apontou dez variáveis associadas à posse de linha telefônica residencial (LTR). Desse conjunto de variáveis, os pesos de pós-estratificação corrigiram os vícios potencias do consumo de feijão em cinco ou mais dias por semana, consumo de folhas, legumes e verduras (FLV) regularmente e recomendado, autoavaliação de saúde considerada ruim e morbidade referida de diabetes,enquanto os pesos eliminaram parcialmente o vício da prevalência de asma, posse de plano de saúde e a cobertura de prevenção de câncer de mama em mulheres de 50 a 59 anos. Conclusões: Tendo em vista a redução dos vícios potencias nos resultados divulgados pelo sistema Vigitel, em regiões de baixa cobertura de telefone fixo, torna-se necessário o uso de procedimentos de ponderação alternativo e a estratégia de seleção de variáveis externas para a construção de pesos de pós-estratificação.

Palavras-chave: Coleta de dados. Entrevista por telefone. Inquéritos

### INTRODUCTION

In 2006, the Department of Health Vigilance of the Ministry of Health implemented the Vigilance System of Chronic Disease Risk Factors by Telephone Survey, known as Vigitel, in the 26 Brazilian capitals and in the Federal District. This system collects information about risk and protection factors pertaining to chronic diseases, such as smoking, eating habits, weight excess, physical exercise, excessive alcohol consumption, and self-reported morbidities, among others<sup>1</sup>. However, the main limitation of Vigitel concerns the exclusion of households without a landline telephone in the northern and northeastern regions. Due to this exclusion, statistical inferences are made by means of mathematical procedures that adjust the results observed in the population with landline telephones to the population without landline telephones. This procedure has the purpose of correcting any potential biases introduced by the limited amount of landline telephones<sup>2,3</sup>.

Although the number of landline telephones has increased starting in 2000 in Brazil, the results of the Survey by Household Samples (*Pesquisa de Amostras por Domicílios* – PNAD)<sup>4</sup> carried out by the Brazilian Institute of Geography and Statistics (IBGE) in 2008 indicate very distinct estimates regarding households served by at least one landline telephone, varying among the capitals from 28% (Boa Vista) to 77% (São Paulo). The northern and northeastern regions have

the lowest percentages – 41 and 45% respectively. In Aracaju, 49% of the households have access to a landline telephone. In the Midwest region, the number reaches 59%, while the highest percentages are in the southeast and South regions, with 70 and 74%, respectively.

Even with this limitation, Vigitel has several appealing features, such as low operational costs and fast publication of results, in comparison to similar household surveys<sup>5-7</sup>. Nevertheless, three issues have been pointed out in scientific productions: the validity of estimates obtained by telephone surveys due to the exclusion of households without a landline telephone<sup>8-11</sup>; the increase in the absence of answers<sup>10,11</sup>; and the methodological procedures used to obtain valid estimates<sup>2,3,12-16</sup>.

Concerning the rate of non-response, Vigitel's experience in Brazil differs from that of other countries, as in the case of the Behavior Risk Factor Surveillance System (BRFSS)<sup>17</sup>, which has been conducted in the United States for over 20 years. While refusals to participate have increased in the BRFSS over the past decade, the refusal rate has been progressively reduced in the case of Vigitel, reaching only 2.5% in its last edition<sup>18</sup>.

Our purpose in this study was to evaluate the effects of using post-stratification to correct biases consequential of the small number of landline telephones in the results divulged by the telephone survey of non-transmissible chronic disease risk and protection factors (Vigitel) conducted in 2008, in Aracaju, Sergipe.

### METHODS

The evaluation of the presence of potential biases in the results of the telephone survey was carried out by comparing the results obtained over the telephone to those of a household survey. The prevalence of the variables of the Household Survey was considered as population values<sup>9</sup>, as they refer to a sample of households with and without landline telephones. To conduct the study, we adopted the following presuppositions: the prevalence of each variable of the Household Survey has insignificant biases; the samples of the Vigitel and of the Household Survey are independent.

The population-based Household Survey interviewed 2,268 adults of both sexes, ranging between 20 and 69 years of age, who lived in 50 census tracts of the city of Aracaju. The survey took place from October 2008 to March 2009; 83% of the interviews were conducted in 2008. The questionnaire applied in the Household Survey is identical to Vigitel's in the modules Feeding Frequency; Smoking; Alcohol; Breast and Cervical Cancer Prevention; Perception; and Self-Reported Morbidity. For the analysis, we selected the variables that compose indicators of food consumption (beans five or more days per week; regular and recommended intake of fruits, greens and vegetables [FGV]; meats with visible fat; whole milk; and sugared soda five or more days per week); of alcohol consumption (abuse and dangerous driving after alcohol abuse); of self-assessment in which health was considered poor; of self-reported morbidities (high blood pressure, diabetes, heart attack, stroke or cerebrovascular accident, high cholesterol or triglycerides, asthma, asthmatic bronchitis,

chronic bronchitis or emphysema, and osteoporosis); of health coverage; of breast cancer prevention in women from 50 to 59 years of age (individuals who had a mammography, and those who had had one in the last two years); and of cervical cancer protection in women from 20 to 59 years of age (had a pap test, and those who had had one in the last three years). These are all qualitative variables, transformed into dichotomous variables that were scored 1 if the answer was positive, and 0 if the answer was negative.

The database of the Household Survey was utilized to identify the possession of landline telephones and to characterize the sociodemographic profile of the population excluded by Vigitel. In the first stage of the study, the adults of the Household Survey were divided in two strata, namely those who had a residential landline telephone (RLT) and those who did not have one. This enabled us to identify the prevalent factors associated with RLT possession by means of hypothesis testing in order to obtain the differences in average between the populations with and without a landline telephone, considering a significance level of 5%. These data also allowed us to characterize the profile of the individuals who did not use a landline telephone through the model of multiple logistic regression  $\left(\log\left\{\frac{\pi(x)}{1-\pi(x)}\right\} = \beta_1 + \beta_2 x_2 + \dots + \beta_p x_p\right)^{19}$ , in which  $\pi(x)$  expresses the probability of lack of access to an RLT according to x<sub>n</sub> characteristics (age range, self-reported skin color, and years of schooling). The explanatory variables were qualitative, and we considered the last category as a reference. The results of a multiple logistic regression are expressed through the odds ratio for a certain x<sub>p</sub> and reference category. An odds ratio of 1 indicates that the odds are equally probable in both groups. A value above 1 indicates how many times the odds are higher in the first group, and a value below 1 indicates how many times the odds are lower in the first than in the second group.

In order to attenuate biases in the estimates of Vigitel due to a limited number of households with landline telephones, the system utilizes post-stratification weights according to age, sex, and schooling, obtained by the propensity scores method<sup>16</sup>. These weights were devised using data from the 2000 Census as an external source, and made available in the database. In order to control the effects of the factor time, new post-stratification weights were devised based on data from the 2008 National Survey by Household Samples. These weights were obtained by means of the ratio between the relative frequency of the population (PNAD, 2008) and of Vigitel in each cell. In total, 36 cells were obtained, composed of sex (F, M), age range ([20-24], [25-34], [35-44], [45-54], [55-64] e [65-69] and schooling years ([0-8], [9-11] e [12-20].

The prevalence proportions  $(\hat{p}_v)$  of Vigitel were obtained by using the sample's basic weight  $\frac{number\_adults}{number\_landline\_telephones}$  multiplied by the post-stratification weights. In turn, the prevalent factors  $(\hat{p}_D)$  of the Household Survey were obtained by using the sample's basic weights and variables of sample planning, as these are data from a complex sampling plan.

The effect of the post-stratification weights on the prevalent factors detected by Vigitel was expressed by the difference between the prevalence divided by the weight's sample, and by prevalence divided by the final weight  $(w_{ij} = \frac{Number of adults_i}{number of landline telephones_i} * post-stratification_weight_i)$  with *i* varying from 1 to n interviews, and *j* varying from 1 to 36 cells.

The definition of bias in an estimate is given by  $[E(\bar{y}) - \bar{Y}]^{1,2}$ , that is, the difference between the expected distribution of the sample's averages and the population value. The lack of knowledge of these parameters is the main difficulty when evaluating biases associated with the sampling process. In this study, we used the difference between two estimates to calculate biases in the Vigitel. Its potential biases are expressed by the difference between the prevalent factors of the Vigitel and of the Household Survey  $(\widehat{bias} (\hat{p}_{vigitel}) = \hat{p}_{vigitel} - \hat{p}_{household})^{8.9}$ . We used the mean squared value (MSE) of the prevalence  $MSE [\hat{p}_{vigitel}] = var [\hat{p}_{vigitel}] + \widehat{bias}^2 (\hat{p}_{vigitel})^{2.3}$  and the square root of the MSE (RMSV) as measures of accuracy. The RMSV informs the distance between the result divulged by the Vigitel system and the population value expressed by estimates of the Household Survey (population with and without landline telephones).

The Household Survey titled "Evaluation of the Effectiveness of the Program Academia da Cidade Aracaju" and the Vigitel were approved by the National Human Research Ethics Committee (CONEP). In the case of Vigitel, the informed consent was replaced by a verbal consent obtained from the interviewees over the telephone.

## RESULTS

Upon comparing the prevalent factors between the individuals who had an RLT to those who did not in the Household Survey, we verified significant differences in 10 of the 22 variables analyzed (Table 1). The group with RLTs was different in prevalence proportions related to the regular and recommended consumption of beans, and fruits, vegetables and greens (FGV); to the consumption of meats with visible fat; to self-assessments in which health was deemed poor; to self-reported morbidities, namely diabetes, asthma, asthmatic bronchitis, and chronic bronchitis or emphysema; to health coverage; and to breast cancer prevention in women from 50 to 59 years of age.

The results of the model of multiple logistic regression, expressed through odds ratios, show that the chance of an adult not being an user of an RLT decreases as schooling years add up. The same relation was found for the age range (Table 2). The variable self-reported skin color was not significant in this model.

In 2008, data from the PNAD<sup>4</sup> indicated that 49% of the households in Aracaju, Sergipe, were equipped with a landline telephone. Due to the exclusion of households without landline telephones, it is possible to observe that the age pyramid of Vigitel's sample is different from the population estimated by PNAD<sup>4</sup> in 2008 if only the sample's basic weights are taken into consideration. This difference also occurred with the variable years of schooling. Upon using propensity procedures to correct statistical inferences, the distribution of Vigitel's sample was adjusted to that of the population estimated by PNAD<sup>4</sup>, which was used as an external source. It is also observable that the use of post-stratification weights in the statistical analyses caused the age pyramid estimated by Vigitel's sample to be equal to that of the population. This also occurred in the distribution of the variable years of schooling (Figure 1).

Variables	With telephone (n <sub>1</sub> )	Without telephone (n <sub>2</sub> )	Difference			
	<b>P</b> <sub>1</sub>	p <sub>2</sub>	<b>p</b> <sub>1</sub> - <b>p</b> <sub>2</sub>	p-value		
Food consumption						
Beans, 5 or more days per week	69.33	80.82	-11.48	0.00		
FGV regularl intake	45.21	30.92	14.30	0.00		
FGV recommended intake	27.12	15.38	11.73	0.00		
Meats with visible fat	27.60	36.47	-8.87	0.00		
Whole milk	52.84	54.92	-2.09	0.54		
Sugared soda, 5 or more days per week	17.18	21.94	-4.76	0.11		
Alcohol consumption						
Abusive	22.51	22.90	-0.39	0.88		
Dangerous driving	5.06	2.53	2.53	0.06		
Smoking						
Smoker	12.04	14.81	-2.77	0.12		
Smokes more than 20 cigarettes per day	4.46	5.16	-0.70	0.60		
Health self-assessment						
Poor health	2.18	4.73	-2.55	0.04		
Self-reported morbidity						
Arterial hypertension	21.60	20.82	0.78	0.71		
Diabetes	5.65	2.75	2.90	0.00		
Heart attack, stroke or cardiovascular accident	5.88	4.13	1.74	0.13		
High cholesterol or triglycerides	2.87	2.65	0.22	0.80		
Asthma, asthmatic bronchitis, chronic bronchitis or emphysema	24.41	13.45	10.97	0.00		
Osteoporosis	4.99	3.17	1.82	0.12		
Health coverage	-					
Has a health plan	48.44	17.77	30.68	0.00		
Breast cancer prevention (50 to 59 years of age) $^{\!\!\!\!\!\!\!^{\scriptscriptstyle\Delta}}$						
Had a mammography exam	91.24	61.66	29.57	0.00		
Had a mammography exam in the last two years	78.93	47.79	31.14	0.00		
Cervical cancer prevention (25 to 59 years of age)§						
Had a pap test	85.25	84.87	0.38	0.89		
Had a pap test in the last three years	78.22	79.47	-1.25	0.66		

Table 1. Prevalent factors according to landline telephone possession detected in the Household Survey. City of Aracaju, Sergipe, 2008.

FGV: fruits, greens and vegetables.

 $(*)n_1=1,088$  (size of the sample with telephone) and  $n_2=1,180$  (size of the sample without telephone).

<sup>6</sup>n, =216 and n,=115.

<sup>s</sup>n<sub>1</sub> = 469 and n<sub>2</sub> = 491.

Upon assessing the effects of the post-stratification weights on the prevalent factors divulged by Vigitel, we found that the prevalence associated with having an RLT presented the biggest differences, except for self-assessments in which health was deemed poor, self-reported diabetes, and mammography examinations. The weights decreased the prevalence of regular and recommended FGV consumption, health plans, and whether female

Table 2. Odds ratio estimates related to non-users of residential landline telephones obtained through the model of multiple logistic regression. Household Survey conducted in the city of Aracaju, Sergipe, 2008.

Variables	Odds ratio	Standard error	t	P>t
Schooling				
0 to 8 years	5.40	1.51	6.02	0.00
9 to 11 years	2.29	0.62	3.04	0.00
12 or more years	1.00			
Age range				
20 to 29 years	4.40	0.77	8.50	0.00
30 to 39 years	3.37	0.65	6.29	0.00
40 to 49 years	2.45	0.52	4.21	0.00
50 or more years	1.00			

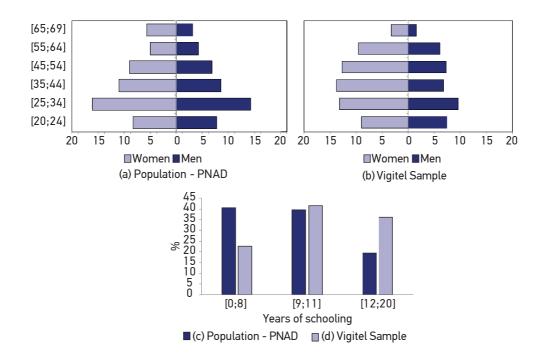


Figure 1. Age pyramid and distribution of the variable years of schooling, according to the surveys. City of Aracaju, Sergipe, 2008.

individuals had had a mammography examination in the last two years, and increased the proportion of consumption of beans five or more days per week, and of meats with visible fat. Among the prevalent factors not associated with having an RLT, the post-stratification weights increased — in less than 2% — the prevalence proportions of consumption of whole milk, and sugared soda and alcohol, while self-reported arterial hypertension, heart attack,

and cervical cancer prevention decreased up to 2.6%. It is also observable that the prevalence proportions that were lower than or equal to 10%, or above 90% presented differences that did not reach 1% (Table 3).

Table 3. Weighted prevalence, standard errors, and absolute differences detected in Vigitel. City of Aracaju, Sergipe, 2008.

	Sample's weight		Final weight			
/ariables	Prevalence (A)	Standard error	Prevalence (B)	Standard error	B-A	
Food consumption						
Beans, 5 or more days per week <sup>#</sup>	68.08	1.22	72.74	1.40	4.66	
FGV regular intake <sup>#</sup>	39.74	1.29	34.27	1.55	5.47	
FGV recommended intake <sup>#</sup>	21.58	1.08	18.48	1.23	3.10	
Meats with visible fat#	24.82	1.15	29.08	1.64	4.25	
Whole milk	46.34	1.32	47.43	1.68	1.09	
Sugared soda, 5 or more days per week	11.56	0.84	12.82	1.15	1.26	
Alcohol consumption						
Abusive	17.75	1.01	19.48	1.39	1.73	
Dangerous driving	2.54	0.42	2.99	0.66	0.45	
Smoking						
Smoker	10.66	0.83	11.50	1.06	0.84	
Smokes more than 20 cigarettes per day	3.70	0.49	4.24	0.71	0.53	
Health self-assessment						
Poor health <sup>#</sup>	2.28	0.39	3.03	0.71	0.74	
Self-reported morbidity						
Arterial hypertension	26.40	1.17	25.20	1.42	1.20	
Diabetes#	6.64	0.65	6.32	0.76	0.32	
Heart attack, stroke or cardiovascular accident	6.67	0.66	5.57	0.62	1.11	
High cholesterol or triglycerides	2.74	0.41	2.90	0.51	0.16	
Asthma, asthmatic bronchitis, chronic bronchitis or emphysema <sup>#</sup>	27.74	1.17	25.05	1.44	2.69	
Osteoporosis	4.61	0.56	4.66	0.70	0.04	
Health coverage						
Has a health plan <sup>#</sup>	54.46	1.33	41.63	1.58	12.82	
Breast cancer prevention (50 to 59 years of age)						
Had a mammography exam <sup>#</sup>	94.74	1.24	94.41	1.67	0.33	
Had a mammography exam in the last two years <sup>#</sup>	81.72	2.34	78.21	3.60	3.51	
Cervical cancer prevention (25 to 59 years of age)						
Had a pap test	85.80	1.41	83.33	1.90	2.46	
Had a pap test in the last three years						

<sup>#</sup>Prevalence related to having a residential landline telephone (RLT). FGV: fruits, greens and vegetables. In the evaluation of potential biases, the results pointed to the presence of absolute biases that were above 2% in 14 of the 22 prevalent factors. Vigitel underestimates the prevalence rates that monitor indicators of food consumption, alcohol consumption, and cervical cancer prevention among women from 25 to 59 years of age. On the other hand, Vigitel overestimates the prevalence rates of self-reported morbidities, the possession of health coverage, and breast cancer prevention among women from 50 to 59 years of age. Out of the 15 prevalent factors, 8 are associated with having an RLT, with biases varying from 2.31 to 12.99% (Table 4).

The post-stratification weights corrected the potential biases of consumption of beans five days or more per week, as well as of regular and recommended FGV intake, and also partially eliminated potential biases regarding the consumption of meats with visible fat, self-reported asthma, health plan possession, and breast cancer prevention among women from 50 to 59 years of age. Among the variables not associated with having an RLT, the presence of biases concerning the consumption of whole milk (7.9%), sugared soda five or more days per week (7.7%), and abusive alcohol intake (5.5%) stand out, as well as self-reported arterial hypertension, at 4.8% (Table 4).

### DISCUSSION

In 2008, only 49% of the permanent, private households located in the capital of Sergipe, Aracaju, were equipped with a landline telephone, according to data from PNAD<sup>4</sup>. Non-users of landline telephones are concentrated in classes with less years of schooling and among younger individuals. These results are consistent with other studies that identified the profile of individuals who do not use landline telephones<sup>10,20</sup>.

The results of the Household Survey point to ten prevalent factors associated with having an RLT, namely regular consumption of beans, regular and recommended consumption of FGV, consumption of meats with visible fat, self-assessment of health deemed poor, self-reported diabetes, asthma, asthmatic bronchitis, chronic bronchitis or emphysema, health coverage, and breast cancer prevention among women between 50 and 59 years old. In this set of prevalent factors, the post-stratification weights corrected potential biases concerning the consumption of beans five or more days per week, and consumption of FGV regularly down to levels below 3%. The weights also partially eliminated the potential bias regarding the prevalence of asthma, possession of a health plan, and breast cancer prevention among women between 50 and 59 years old. Bernal<sup>21</sup> found a similar result for the prevalence of asthma in a study conducted in the city of Rio Branco, Acre. However, in Aracaju, we observed that the method of propensity scores had a poor performance in reducing the biases in prevalent factors that were below 10% or above 90%, detected in the population with landline telephones, as it occurred in the case of mammography exams among women between 50 and 69 years old; in the population with landline telephones, the percentage reaches 95%. It was expected that the post-stratification weights would decrease this prevalence, as the rate of mammography exams in this population is 83%.

	Household Survey	Vigitel					
Variables		Prevalence	Variance	Bias	MSE	RMSE	
		p <sub>v</sub>		$p_v - p_D$	Var + Bias²	$\sqrt{MSE}$	
Food consumption							
Beans, 5 or more days per week <sup>#</sup>	75.06	72.74	1.97	-2.31	7.31	2.70	
FGV regularl intake <sup>#</sup>	36.46	34.27	2.41	-2.19	7.22	2.69	
FGV recommended intake <sup>#</sup>	19.46	18.48	1.52	-0.98	2.47	1.57	
Meats with visible fat <sup>#</sup>	33.78	29.08	2.69	-4.71	24.84	4.98	
Whole milk	55.30	47.43	2.82	-7.86	64.65	8.04	
Sugared soda, 5 or more days per week	20.43	12.82	1.33	-7.61	59.27	7.70	
Alcohol consumption							
Abusive	24.81	19.48	1.93	-5.33	30.35	5.51	
Dangerous driving	3.89	2.99	0.44	-0.89	1.24	1.11	
Smoking							
Smoker	13.12	11.50	1.12	-1.63	3.76	1.94	
Smokes more than 20 cigarettes per day	4.59	4.24	0.51	-0.35	0.63	0.79	
Health self-assessment							
Poor health <sup>#</sup>	3.18	3.03	0.51	-0.16	0.53	0.73	
Self-reported morbidity							
Arterial hypertension	20.57	25.20	2.01	4.63	23.47	4.84	
Diabetes <sup>#</sup>	3.77	6.32	0.58	2.55	7.08	2.66	
Heart attack, stroke or cardiovascular accident	4.63	5.57	0.38	0.93	1.26	1.12	
High cholesterol or triglycerides	2.63	2.90	0.26	0.27	0.33	0.57	
Asthma, asthmatic bronchitis, chronic bronchitis or emphysema <sup>#</sup>	17.34	25.05	2.08	7.71	61.56	7.85	
Osteoporosis	3.73	4.66	0.50	0.93	1.35	1.16	
Health coverage							
Has a health plan <sup>#</sup>	32.01	41.63	2.48	9.63	95.16	9.76	
Breast cancer prevention (50 to 59 years o	f age)						
Had a mammography exam <sup>#</sup>	82.50	94.41	2.79	11.91	144.65	12.03	
Had a mammography exam in the last two years <sup>#</sup>	67.24	78.21	12.99	10.97	133.37	11.55	
Cervical cancer prevention (25 to 59 years of age)							
Had a pap test	86.39	83.33	3.60	-3.06	12.95	3.60	
Had a pap test in the last three years	80.62	78.05	4.26	-2.57	10.85	3.29	

Table 4. Mean squared error of the prevalent factors detected in Vigitel. City of Aracaju, Sergipe, 2008.

\*Prevalence related to having a residential landline telephone (RLT). FGV: fruits, greens and vegetables. However, the international scientific production on the presence of biases in Telephone Surveys<sup>8-11</sup> shows that propensity scores eliminate biases only partially. The results found in the Vigitel Aracaju system corroborate the international literature.

With the purpose of reducing potential biases, another strand of research uses alternative methods known as *dual frame*<sup>22</sup> and *multiple mode and frame*<sup>23</sup> surveys. These methods utilize more than one set of records for random sampling and have some advantages over the Telephone Survey. Previously, the greatest challenge was obtaining valid estimates for data from multiple sets of records, but with current technological advancements, algorithms are available to aid in obtaining valid estimates. Nevertheless, Flores-Cervantes et al.<sup>13</sup> alert that, for surveys with a low rate of households with landline telephones, the dual frame survey has a relatively high cost, which also makes its use unfeasible.

### CONCLUSION

In Aracaju, Sergipe, non-users of landline telephones are concentrated in less favored social classes, and present a higher prevalence in variables related to risk factors, such as the consumption of meat with visible fat.

In general, the post-stratification weights partially eliminated estimate biases, whose origin is associated with the limited number of households equipped with a landline telephone. However, in Aracaju, we observed that the method of propensity scores had a poor performance in reducing biases when a given prevalence detected in the population with landline telephones was below 10% or above 90%.

With the purpose of reducing potential biases in results divulged by the Vigitel system in regions with a small number of landline telephones, it is necessary to use alternative weighting procedures and strategies of selection of external variables in order to devise post-stratification weights. The use of alternative methods with multiple sets of records is also necessary in order to increase Vigitel's coverage.

In Brazil, the challenge lies in the lack of methodological studies on the estimation of biases resulting from the absence of response and from the limited sets of records that are used in random sampling.

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